

# Cryptography Foundations

## Solution Exercise 8

### 8.1 Changing the Distribution of Bit-Guessing Problems

- a) Recall the solution to Exercise 1.3 a): We defined the set  $\mathcal{X}^* := \{x \in \mathcal{X} \mid P_X(x) \geq P_Y(x)\}$  and proved that  $\delta(X, X') = \Pr[X \in \mathcal{X}^*] - \Pr[X' \in \mathcal{X}^*]$ . Stated differently,  $\mathcal{X}^*$  is the set of elementary events and  $\mathcal{X}^* \subseteq \mathcal{X}$  is a particular event and we write  $\delta(X, X') = \Pr^X[\mathcal{X}^*] - \Pr^{X'}[\mathcal{X}^*]$ . As explained in that solution, the maximum likelihood method allowed us to derive this event  $\mathcal{X}^*$  which maximizes the term  $\Pr^X[\mathcal{X}^*] - \Pr^{X'}[\mathcal{X}^*]$ . Briefly, adding an additional element  $x \in \mathcal{X} \setminus \mathcal{X}^*$  to the set  $\mathcal{X}^*$  would only decrease the term (because  $P_X(x) - P_{X'}(x) < 0$ ) and also removing an element from the set  $\mathcal{X}^*$  would only decrease the term (because  $P_X(x) - P_{X'}(x) \geq 0$ ).

More generally, for two probability spaces  $(\Omega, \mathcal{F}, P_X)$  and  $(\Omega, \mathcal{F}, P_{X'})$  we have that

$$\delta(X, X') = \sup_{\mathcal{B} \in \mathcal{F}} \left| \Pr^X[\mathcal{B}] - \Pr^{X'}[\mathcal{B}] \right|,$$

which is another common formulation of the statistical distance. (The text above actually is the proof for the special case in which we have the (finite) sample space  $\Omega = \mathcal{X}$  and event set  $\mathcal{F} = 2^\Omega$ , i.e., full information. This is the typical case in this lecture.)

It is not hard to see that for any event  $\mathcal{A} \in \mathcal{F}$  we have

$$\Pr^X[\mathcal{A}] - \Pr^{X'}[\mathcal{A}] \leq \sup_{\mathcal{B} \in \mathcal{F}} \left( \Pr^X[\mathcal{B}] - \Pr^{X'}[\mathcal{B}] \right) \leq \sup_{\mathcal{B} \in \mathcal{F}} \left| \Pr^X[\mathcal{B}] - \Pr^{X'}[\mathcal{B}] \right| = \delta(X, X').$$

- b) Exercise 4.4 in the lecture notes asks to show that for a bit-guessing problem  $(S, B)$  and a distinguisher  $D$  for it, if one changes the instance distribution of  $(S, B)$  by at most  $d$  in terms of statistical distance, then the performance of  $D$  changes by at most  $2d$ . The performance of  $D$  is measured in terms of its advantage  $\Lambda^D((S, B))$ . Changing the instance distribution of  $(S, B)$  as described above means considering a new bit-guessing problem  $(S', B')$  such that  $d = \delta((S, B), (S', B'))$ . We assume without loss of generality that the output bit  $B$  of  $S$  is a deterministic function of  $S$  and thus the statistical distance of  $\delta((S, f(S)), (S', f(S')))$  is no greater than  $\delta(S, S')$  as we know from a previous exercise. In summary: what we want to prove is in this case

$$\Lambda^D((S, B)) = \Lambda^D((S', B')) + 2 \cdot \delta(S, S').$$

Consider the random experiment  $D(S, B)$ , i.e., a distinguisher  $D$  interacting with system  $S$  (which outputs bit  $B$ ) and outputs a guess  $Z$ , as a probability space where the elementary events correspond to sampling  $D$  and sampling  $S$ . All properties, including the event  $\mathcal{A} := Z = B$  are deterministic functions when given these (sampled) problem instance and distinguisher. From subtask a), we conclude that

$$\begin{aligned} \Lambda^D((S, B)) - \Lambda^D((S', B')) &= 2 \cdot \Pr^{D(S, B)}[Z = B] - 1 - (2 \cdot \Pr^{D(S', B')}[Z = B'] - 1) \\ &= 2 \cdot (\Pr^{D(S, B)}[\mathcal{A}] - \Pr^{D(S', B')}[\mathcal{A}]) \\ &\leq 2 \cdot \delta((D, S), (D, S')) \leq 2 \cdot \delta(S, S'). \end{aligned}$$

Note that  $Z = B$  and  $Z' = B'$  denote the same event in the two experiments (expressed as a function of  $D$  and  $S$ )<sup>1</sup>. The final step that  $\delta((D, S), (D, S')) \leq \delta(S, S')$  follows from a simple property of the statistical distance (analog to one of the properties proven on the previous exercise sheet) since by definition of the random experiment,  $D$  and  $S$  (resp.  $S'$ ) are sampled independently.

## 8.2 Amplifying the Performance of a Worst-Case Solver

Let  $X_i$  for  $i \in \{1, \dots, q\}$  be the binary random variable that is 1 if the  $i$ th invocation of  $S$  returns the correct bit. Since  $S$  has performance  $\epsilon$ , we have  $p := \Pr[X_i = 1] = \frac{\epsilon}{2} + \frac{1}{2}$ . Note that all  $X_i$  are independent and that the solver  $T$  outputs the wrong bit if and only if  $S$  outputs more wrong than correct bits. That is, the probability that  $T$  outputs the wrong bit is  $\Pr\left[\sum_{i=1}^q X_i < \frac{q}{2}\right]$ . Let  $\alpha := \frac{\epsilon}{2} = p - \frac{1}{2}$ . We then obtain for the probability that  $T$  outputs the wrong bit using Hoeffding's inequality

$$\Pr\left[\sum_{i=1}^q X_i < \frac{q}{2}\right] = \Pr\left[\sum_{i=1}^q X_i \leq (p - \alpha)q\right] \leq e^{-2\alpha^2 q} = e^{-q\epsilon^2/2}.$$

For  $q \geq \frac{2}{\epsilon^2} \cdot \log \frac{2}{\delta}$ , we have

$$e^{-q\epsilon^2/2} \leq e^{-\log(2/\delta)} = e^{\log(\delta/2)} = \frac{\delta}{2}.$$

Hence, the success probability of  $T$  for such  $q$  is at least  $1 - \frac{\delta}{2}$ , and the performance of  $T$  is at least  $1 - \delta$ .

## 8.3 The Next Bit Test

Recall that for an integer  $i$  the notation  $a^i$  denotes the sequence  $a_1, \dots, a_i$ , and that we denote its concatenation with another sequence  $b^j$  (namely, the sequence  $a_1, \dots, a_i, b_1, \dots, b_j$ ) as  $a^i b^j$ . For this task we further introduce the following notation: for integers  $i \leq j$ , we write  $a^{i:j}$  to denote the sequence  $a_i, a_{i+1}, \dots, a_j$  (note that  $a^{j:i}$  would correspond to the empty sequence). We now describe how to construct a predictor  $P_i$ , with  $i \in \{1, \dots, \ell\}$ , for the  $i$ -th bit of an arbitrarily distributed bit-string  $X^\ell$ . First,  $P_i$  receives the (partial) bit-string  $X^{i-1}$ . Then it samples the bit-string  $U^{i:\ell}$  uniformly at random (i.e., each bit  $U_i, \dots, U_\ell$  is distributed independently and uniformly at random).  $P_i$  then proceeds by invoking  $D$  on input the bit-string  $X^{i-1} U^{i:\ell}$ . Upon  $D$  outputting a guess bit  $Z$ ,  $P_i$  outputs as its guess for  $X_i$  the bit  $Z \oplus U_i$ .

Before analyzing the advantage of the predictor  $P_i$ , let introduce the following hybrid sequences:

$$\mathbf{H}_k := X^k U^{k+1:\ell} \tag{1}$$

Note that for the extreme cases we have

$$\mathbf{H}_0 = U^\ell \quad \text{and} \quad \mathbf{H}_\ell = X^\ell. \tag{2}$$

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<sup>1</sup>This means that we can identify the subset of pairs of deterministic systems from the product space  $\mathcal{D} \times \mathcal{S}$  for which the output bit of the distinguisher equals the bit of the bit-guessing problem.

Then for any  $i \in \{1, \dots, \ell\}$  we have:

$$\begin{aligned}
\Lambda^{P_i}((X^{i-1}, X_i)) &= 2 \cdot \Pr^{P_i(X^{i-1}, X_i)}[Z' = X_i] - 1 \\
&= 2 \cdot \Pr^{P_i(X^{i-1}, X_i)}[Z \oplus U_i = X_i] - 1 \\
&= 2 \cdot \left( \Pr^{P_i(X^{i-1}, X_i)}[Z = X_i \oplus U_i \mid U_i = X_i] \cdot \frac{1}{2} \right. \\
&\quad \left. + \Pr^{P_i(X^{i-1}, X_i)}[Z = X_i \oplus U_i \mid U_i \neq X_i] \cdot \frac{1}{2} \right) - 1 \\
&= \Pr^{D(X^{i-1} X_i U^{i+1:\ell})}[Z = 0] + \Pr^{D(X^{i-1} \bar{X}_i U^{i+1:\ell})}[Z = 1] - 1 \\
&= \Pr^{D(X^{i-1} \bar{X}_i U^{i+1:\ell})}[Z = 1] - \Pr^{D(X^{i-1} X_i U^{i+1:\ell})}[Z = 1] \\
&= \Delta^D(X^{i-1} X_i U^{i+1:\ell}, X^{i-1} \bar{X}_i U^{i+1:\ell}).
\end{aligned}$$

Now consider a (probabilistic) system  $S$  which outputs the sequence  $X^{i-1} U^{i+1:\ell}$ . Recall from Exercise 1.3 b) that for a bit  $B$  correlated with  $S$  and an independent and uniformly distributed bit  $U$ , we have

$$\Delta^D((S, B), (S, U)) = \frac{1}{2} \cdot \Delta((S, B), (S, \bar{B})). \quad (3)$$

Therefore, since  $X_i$  is indeed correlated with  $S$ , whereas  $U_i$  is independent and uniformly distributed, from (3) we get

$$\Delta^D(X^{i-1} X_i U^{i+1:\ell}, X^{i-1} \bar{X}_i U^{i+1:\ell}) = 2 \cdot \Delta^D(X^i U^{i+1:\ell}, X^{i-1} U^{i:\ell}).$$

Putting things together, using (1) we have

$$\Lambda^{P_i}((X^{i-1}, X_i)) = 2 \cdot \Delta^D(\mathbf{H}_i, \mathbf{H}_{i-1}).$$

Finally, using (a slight variation of) Lemma 2.2 and (2), we have

$$\sum_{i=1}^{\ell} \Lambda^{P_i}((X^{i-1}, X_i)) = 2 \cdot \sum_{i=1}^{\ell} \Delta^D(\mathbf{H}_i, \mathbf{H}_{i-1}) = 2 \cdot \Delta^D(\mathbf{H}_\ell, \mathbf{H}_0) = 2 \cdot \Delta^D(X^\ell, U^\ell),$$

and thus it follows that not all predictors  $P_i$  can have advantage less than  $\frac{2}{\ell} \cdot \Delta^D(X^\ell, U^\ell)$ . Turned around, this means that there exists an  $i \in \{1, \dots, \ell\}$  and a predictor  $P_i$  for  $X^\ell$  such that

$$\Lambda^{P_i}((X^{i-1}, X_i)) \geq \frac{2}{\ell} \cdot \Delta^D(X^\ell, U^\ell).$$